The m-spread model using branching processes

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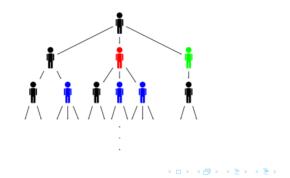
Joint work with J.C. Ban and Y.L Wu

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Motivation

When considering the evolution of a population, a branching process is often a good tool to describe it.

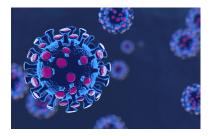
In a classical Galton-Watson branching process, the production of each individual is assumed to be independent of the past of the population.



Motivation

In reality, for example, when modeling the spread of a certain virus, the infection is more complicated.

The infectious period of the carriers usually last a while and before the recovery the individuals are still capable to pass on the viruses.



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Motivation

Goals:

- 1. Construct a probability model to describe the evolution when the reproduction mechanism shows some dependence of the past of the population.
- 2. Study what happens to this kind of population in a long run.



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Outline







The induced branching process from the m-spread model

Outline



2 Main results

In the induced branching process from the m-spread model



Assumptions

Assume that all the individuals of the population are categorized into k different types.

Assume that the infectivity of each individual after infected will last for some time, say m units of time, and, during the infectious period, the individual has ability to pass the viruses to others.

Therefore, the type composition of the next generation depends on the type structure of the past m generations in the spread history.



Let $\mathcal{A} = \{a_1, a_2, \cdots, a_k\}$ be the type set.

For $d \in \mathbb{N}$, we use the following notations:

- 1. T_d is the conventional *d*-tree.
- 2. |g| stands for the length of $g \in T_d$, i.e., the number of the edges from the root ϵ of T_d to g.
- 3. Define $\Delta_n = \{g \in T_d : 0 \le |g| \le n\}.$
- 4. Define $\Delta_m^n = \Delta_n \setminus \Delta_m = \{g \in T_d : m < |g| \le n\}.$
- 5. Let $F \subseteq \Delta_m$ and assume at least one $g \in F$ such that |g| = m. Every function

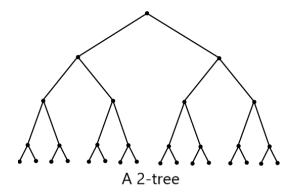
$$\alpha: F \to \mathcal{A}$$

is called an *m*-order pattern (or *m*-pattern for brevity).

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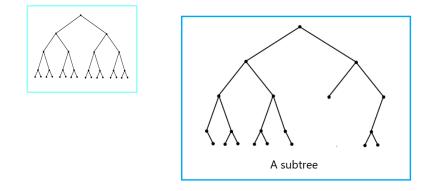
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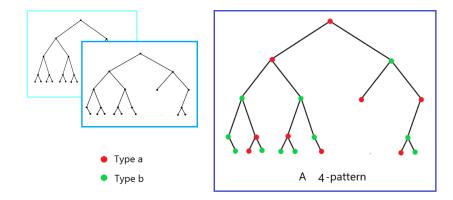


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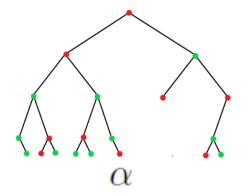
Let \mathcal{P}_n be the set of the collection of all the *n*-patterns. For each *n*-pattern $\alpha \in \mathcal{P}_n$, we define the following:

- (i) The parent $\alpha^{(0)}$ of α , which is the subtree from the root of α up to the (n-1)th level in α .
- (ii) For each $r = 1, 2, \dots, n$, let $\Sigma_{\alpha}^{(r)}$ be the set of all the type-attached vertices in the *r*th level in α .
- (iii) Let $\sigma_{\alpha}^{(r)} = |\Sigma_{\alpha}^{(r)}|$ be the number of vertices in $\Sigma_{\alpha}^{(r)}$. So, we can write

$$\Sigma_{\alpha}^{(r)} = \{g_1^{\alpha,r}, g_2^{\alpha,r}, \cdots, g_{\sigma_{\alpha}^{(r)}}^{\alpha,r}\}.$$

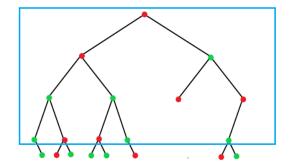
(iv) For each $r = 1, 2, \dots, n-1, j = 1, 2, \dots, \sigma_{\alpha}^{(r)}$, let $\alpha_j^{(r)}$ be the subtree rooted at $g_j^{\alpha,r}$ in the *r*th level of α and up to the *n*th level in α . We call $\alpha_j^{(1)}$ the *j*th "son" of α and $\alpha_j^{(r)}$ the *j*th "*r*-descendant" of α .

For example, d = 2, n = 4 and $\mathcal{A} = \{a, b\}$.



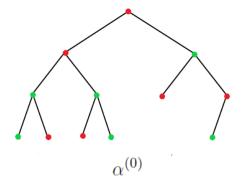
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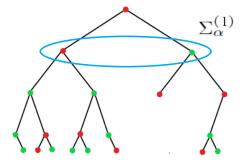


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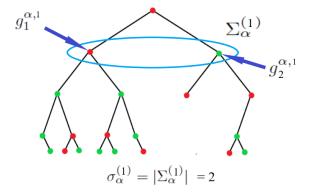
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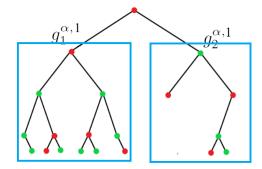
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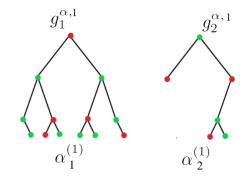
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Let \mathcal{Q}_m be a subset of \mathcal{P}_m such that, for every $\beta \in \mathcal{Q}_m$, there exist $q \in \mathcal{Q}_m$ such that

$$\beta_j^{(1)} = q^{(0)}$$
 for all $j = 1, 2, \cdots, \sigma_\beta^{(1)}$.

Since \mathcal{P}_m is finite, so is \mathcal{Q}_m .

Let $\mathcal{Q}_m^{(0)} = \{\beta^{(0)} : \beta \in \mathcal{Q}_m\}$ be the collection of the parents of all *m*-patterns in \mathcal{Q}_m . We call $\mathcal{Q}_m^{(0)}$ the parent set of \mathcal{Q}_m .

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Note that $\mathcal{Q}_m^{(0)} \subseteq \mathcal{P}_{m-1}$.

Let $f: \mathcal{Q}_m^{(0)} \times \mathcal{Q}_m \to [0, 1]$ be a function such that (i) for any $\alpha \in \mathcal{Q}_m^{(0)}$ and any $\beta \in \mathcal{Q}_m$, if $\beta^{(0)} \neq \alpha$, then $f(\alpha, \beta) = 0$; (ii) for each $\alpha \in \mathcal{Q}_m^{(0)}$, $\sum_{\beta \in \mathcal{Q}_m} f(\alpha, \beta) = 1$.

So, for any fixed $\alpha \in \mathcal{Q}_m^{(0)}$, $f(\alpha, \cdot) \equiv \{f(\alpha, \beta) : \beta \in \mathcal{Q}_m\}$ forms a probability distribution.

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Each $f(\alpha, \beta)$ can be considered as the probability for the (m-1)-pattern α to grow into the *m*-pattern β .

The m-spread model

For example, d = 2, $\mathcal{A} = \{a, b\}$, m = 3, then

$$f: \mathcal{Q}_3^{(0)} \times \mathcal{Q}_3 \to [0,1]$$

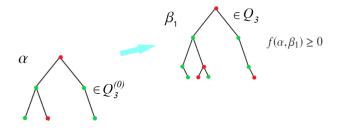
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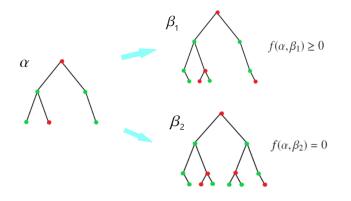
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Therefore, by Kolmogorov extension theorem, we can construct a probability space $(\Omega, \mathcal{F}, \mathbb{P})$ and random elements

$$\{\zeta_{\alpha}: \alpha \in \mathcal{Q}_m^{(0)}\}$$

such that, for each α , ζ_{α} is a \mathcal{Q}_m -valued random element on $(\Omega, \mathcal{F}, \mathbb{P})$ such that $(\zeta_{\alpha})^{(0)}(\omega) = \alpha$ for each $\omega \in \Omega$ and the probability of ζ_{α} taking value at $\beta \in \mathcal{Q}_m$ is $f(\alpha, \beta)$.

The collection $\mathcal{R} = \{\zeta_{\alpha} : \alpha \in \mathcal{Q}_m^{(0)}\}$ of random *m*-patterns is called the *random m-spread model* over the type set \mathcal{A} with pattern distribution f on $\mathcal{Q}_m^{(0)} \times \mathcal{Q}_m$.

Note that, if $\zeta_{\alpha} \in \mathcal{R}$ and $f(\alpha, \beta) > 0$ for $\beta \in \mathcal{Q}_m$, then

$$\zeta_{\beta_j^{(1)}} \in \mathcal{R} \qquad \text{for all } j = 1, 2, \cdots, \sigma_{\beta}^{(1)}.$$

Assume that the first m generations of the population have existed initially and the type structure is α .

Let $\tau_{\alpha}^{m-1} = \alpha$, for some $\alpha \in \mathcal{Q}_m^{(0)}$ represent the initial type structure for the population from time (i.e., generation) 0 up to time (i.e., generation) m-1.

We will construct a sequence $\{\tau_{\alpha}^{m+n}\}_{n\geq 0}$ of random elements on the probability space $(\Omega, \mathcal{F}, \mathbb{P})$ and each realization of $\{\tau_{\alpha}^{m+n}\}_{n\geq 0}$ can be considered as a growing tree and can be used to describe the evolution of the type structure of the population over time.

Since $\alpha \in \mathcal{Q}_m^{(0)}$, it is the parent of some *m*-pattern in \mathcal{Q}_m .

So, at time 0, we can replace α with an *m*-pattern $\alpha_i^m \in \mathcal{Q}_m \equiv \{\alpha_i^m : 1 \leq i \leq l_m\}$ to obtain the value of the random element τ_{α}^m on $(\Omega, \mathcal{F}, \mathbb{P})$ with probability

$$\mathbb{P}(\tau^m_\alpha=\alpha^m_i)=f(\alpha,\alpha^m_i)$$

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for all $i = 1, 2, \dots, l_m$.

Note that τ^m_{α} has the same distribution as the random *m*-pattern ζ_{α} .

We assume that the replacements of each (m-1)-pattern with an *m*-pattern is independent of the simultaneous replacement of other (m-1)-patterns.

Since, if $\tau_{\alpha}^{m} = \alpha_{i_{0}}^{m}$, then $(\alpha_{i_{0}}^{m})_{j}^{(1)} \in \mathcal{Q}_{m}^{(0)}$ for $j = 1, 2, \cdots, \sigma_{\alpha_{i_{0}}^{m}}^{(1)}$. So, at time 1, we replace $(\alpha_{i_{0}}^{m})_{j}^{(1)}$ with some $\alpha_{i_{j}}^{m} \in \mathcal{Q}_{m}$ to obtain a (m+1)-pattern $\alpha_{i}^{m+1} \in \mathcal{P}_{m+1}$ as the value of τ_{α}^{m+1} with probability

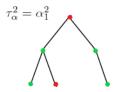
$$\mathbb{P}(\tau_{\alpha}^{m+1} = \alpha_{i}^{m+1} | \tau_{\alpha}^{m} = \alpha_{i_{0}}^{m})$$

$$= \prod_{j=1}^{\sigma_{\alpha_{i_{0}}^{m}}} \mathbb{P}((\tau_{\alpha}^{m+1})_{j}^{(1)} = \alpha_{i_{j}}^{m} | \tau_{\alpha}^{m} = \alpha_{i_{0}}^{m}) = \prod_{j=1}^{\sigma_{\alpha_{i_{0}}^{m}}} f((\alpha_{i_{0}}^{m})_{j}^{(1)}, \alpha_{i_{j}}^{m}).$$

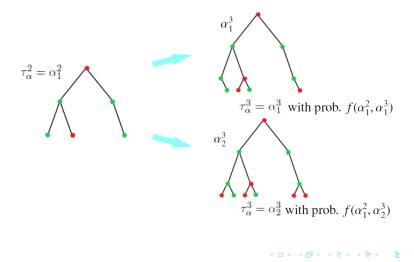
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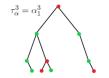
For example, d = 2, $\mathcal{A} = \{a, b\}$, m = 3. Assume that initially we have $\alpha = \alpha_1^2 \in \mathcal{Q}_3^{(0)}$.



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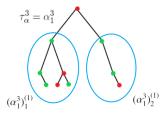
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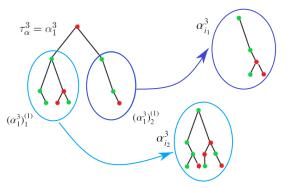
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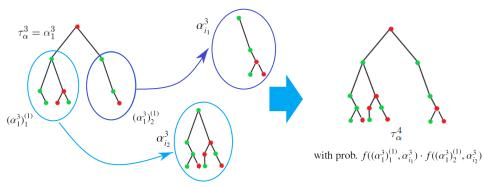
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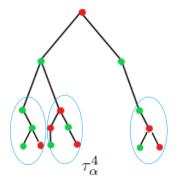


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It should be noted that the type structure in the first m + 1 generations of τ_{α}^{m} and τ_{α}^{m+1} are identical with probability 1.

Therefore, we can consider that τ_{α}^{m} is growing into τ_{α}^{m+1} . Here, τ_{α}^{m+1} is a random element taking values in \mathcal{P}_{m+1} .

Assume that τ_{α}^{m+n} is constructed and note that τ_{α}^{m+n} takes values in \mathcal{P}_{m+n} .

Given $\tau_{\alpha}^{m+n} = \alpha_{i_0}^{m+n} \in \mathcal{P}_{m+n}$, we then replace each (m-1)-pattern $(\alpha_{i_0}^{m+n})_j^{(n+1)} \in \mathcal{Q}_m^{(0)}$ which is rooted at $g_j^{\alpha_{i_0}^{m+n}, n+1}$, for $j = 1, 2, \cdots, \sigma_{\alpha_{i_0}^{m+n}}^{(n)}$,

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with some $\alpha_{i_j}^m \in \mathcal{Q}_m$ to obtain a (m+n+1)-pattern α_i^{m+n+1} as the value of τ_{α}^{m+n+1} with probability

$$\mathbb{P}(\tau_{\alpha}^{m+n+1} = \alpha_{i}^{m+n+1} | \tau_{\alpha}^{m+n} = \alpha_{i_{0}}^{m+n})$$

$$= \prod_{\substack{j=1\\ \sigma_{\alpha_{i_{0}}}^{(n+1)}}} \mathbb{P}((\tau_{\alpha}^{m+n+1})_{j}^{(n+1)} = \alpha_{i_{j}}^{m} | \tau_{\alpha}^{m+n} = \alpha_{i_{0}}^{m+n})$$

$$= \prod_{\substack{j=1\\ j=1}} f((\alpha_{i_{0}}^{m+n})_{j}^{(n+1)}, \alpha_{i_{j}}^{m}),$$

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We continue this process and obtain a sequence $\{\tau_{\alpha}^{m+n}\}_{n\geq 0}$ of almost surely growing random trees.

Therefore, the limit $\tau_{\alpha} \equiv \lim_{n \to \infty} \tau_{\alpha}^{m+n}$ exists with probability 1.

The random element
$$\tau_{\alpha}$$
 takes values on $\mathcal{A}^{\bar{d}}$, where
 $\bar{d} = \max_{\beta \in \mathcal{Q}_m} (\max_{0 \le i \le m} \sigma_{\beta}^{(i)}).$

For almost every $\omega \in \Omega$, $\tau_{\alpha}^{m+n}(\omega)$ is the subtree from the root of $\tau_{\alpha}(\omega)$ to the (m+n)th level of $\tau_{\alpha}(\omega)$, for all $n = 0, 1, 2, \cdots$.

We call τ_{α} the infinite random spread pattern induced from the random m-pattern ζ_{α} with respect to the random m-spread model \mathcal{R} and pattern distribution f.

Question: If we go to the *n*th level of the infinite tree τ_{α} , what is the proposition of a given type a_i as $n \to \infty$?

More generally, if $\mathcal{O}_a(\tau_{\alpha}|_{\Delta_r^s(\tau_{\alpha})})$ is the occurrences of $a \in \mathcal{A}$ in τ_{α} from level r+1 to level s and $\{k_n\}_{n=1}^{\infty}$ is an increasing sequence such that $k_n \to \infty$ as $n \to \infty$, then what happens to the rate:

where
$$s_{\alpha}(a; \{k_n\}_{n=1}^{\infty})$$

$$:= \lim_{n \to \infty} s_{\alpha}(a; [s_n, s_{n+1}]) := \lim_{n \to \infty} \frac{\mathcal{O}_a(\tau_{\alpha}|_{\Delta_r^s(\tau_{\alpha})})}{|\Delta_r^s(\tau_{\alpha})|}$$
where $s_n = \sum_{i=1}^n k_i$.

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Outline





The induced branching process from the *m*-spread model



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Main results

Recall that τ_{α} is the infinite spread pattern induced from the random *m*-pattern ζ_{α} with respect to the random *m*-spread model \mathcal{R} and the type set $\mathcal{A} = \{a_i\}_{i=1}^k$.

Let
$$\mathbf{A} \equiv \mathcal{Q}_m^{(0)} = \{\alpha_i\}_{i=1}^{\mathbf{k}}$$
.

For each $a \in \mathcal{A}$, we define

$$\theta(a) = \{\beta \in \mathbf{A} : \epsilon(\beta) = a\}$$

where $\epsilon(\beta)$ is the root of the pattern β .

Main results

Theorem 1 (Ban, Hong and Wu, 2022+).

Let τ_{α} be the infinite spread pattern with the type set $\mathcal{A} = \{a_i\}_{i=1}^k$. Suppose that $\{k_n\}_{n=1}^{\infty}$ is a sequence of positive integers such that $k_n \to \infty$ as $n \to \infty$. Let $s_n = \sum_{r=1}^n k_n$. Then, for any $a_i \in \mathcal{A}$, we have that

$$s_{\alpha}(a_j; \{k_n\}_{n=1}^{\infty}) = \lim_{n \to \infty} \frac{\mathcal{O}_{a_j}(\tau_{\alpha}|_{\Delta_{s_n}^{s_{n+1}}(\tau_{\alpha})})}{\left|\Delta_{s_n}^{s_{n+1}}(\tau_{\alpha})\right|} = \sum_{\substack{i=1,2,\cdots,\mathbf{k} \ s.t.\\\alpha_i \in \theta(a_j)}} \mathbf{v}_i$$

with probability 1, where $\vec{\mathbf{v}} = (\mathbf{v}_1, \cdots, \mathbf{v}_k)$ is the left eigenvector of the offspring mean matrix \mathbf{M} of the induced branching process $\{\vec{\mathbf{Z}}_n\}_{n=0}^{\infty}$ associated with its maximal eigenvalue.

Main results

Theorem 2 (Ban, Hong and Wu, 2022+).

Under the hypotheses of Theorem 1, if $k_n = kn$ for all $n \in \mathbb{N}$, then, we have that

$$s_{\alpha}(a_j;k) \equiv \lim_{n \to \infty} \frac{\mathcal{O}_{a_j}(\tau_{\alpha}|_{\Delta_{kn}^{k(n+1)}(\tau_{\alpha})})}{\left|\Delta_{kn}^{k(n+1)}(\tau_{\alpha})\right|} = \sum_{\substack{i=1,2,\cdots,\mathbf{k} \ s.t.\\\alpha_i \in \theta(a_j)}} \mathbf{v}_i$$

with probability 1.

In particular, if k = 1, then $\mathcal{O}_a(\tau_\alpha|_{\Delta_{n-1}^n(\tau_\alpha)})$ is the occurrences of the type a in the *n*th generation of τ_α .

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The induced branching process from the m-spread model

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Outline

① The *m*-spread model

2 Main results

3

The induced branching process from the m-spread model

We will construct a multitype branching process $\{\vec{\mathbf{Z}}_n\}_{n\geq 0}$ on the same probability space $(\Omega, \mathcal{F}, \mathbb{P})$ in which each (m-1)-pattern is considered as a type.

Let
$$\mathbf{A} = \mathcal{Q}_m^{(0)} = \{\alpha_1, \alpha_2, \cdots, \alpha_k\}.$$

If, in the original population, $\tau_{\alpha}^{m-1} = \alpha = \alpha_{i_0}$ for some $\alpha_{i_0} \in \mathbf{A}$, then we let the induced population $\{\vec{\mathbf{Z}}_n\}_{n\geq 0}$ start with an individual of type $\alpha_{i_0} \in \mathbf{A}$ at time 0, i.e., $\vec{\mathbf{Z}}_0 = \vec{e}_{i_0}$, where \vec{e}_{i_0} is the standard unit vector with 1 as its i_0 th component and 0 elsewhere.

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We also write
$$\{\vec{\mathbf{Z}}_n^{(i_0)}\}_{n\geq 0}$$
 for $\{\vec{\mathbf{Z}}_n\}_{n\geq 0}$ when $\vec{\mathbf{Z}}_0 = \vec{e}_{i_0}$.

After one unit of time, when the type structure of the population in the *m*-spread model grows from τ_{α}^{m-1} to τ_{α}^{m} , at the same time, in the induced population, we replace the initial (m-1)-pattern α_{i_0} with $\sigma_{\tau_{\alpha}}^{(1)}$, the (m-1)-patterns rooted at level 1 in τ_{α}^{m} , to obtain the population

$$\left\{ (\tau_{\alpha}^{m})_{1}^{(1)}, (\tau_{\alpha}^{m})_{2}^{(1)}, \cdots, (\tau_{\alpha}^{m})_{\sigma_{\tau_{\alpha}}^{(1)}}^{(1)} \right\}$$

at time 1, which are called the *children* of the initial (m-1)-pattern in the language of branching process.

Note that each $(\tau_{\alpha}^{m})_{i}^{(1)}$ is a random element taking values in the type set **A**.

Let $\mathbf{Z}_{1,i}^{(i_0)}$ be the number of the (m-1)-patterns of type α_i in the population at time 1, $i = 1, 2, \cdots, \mathbf{k}$, and call the vector $\vec{\mathbf{Z}}_1^{(i_0)} = (\mathbf{Z}_{1,1}^{(i_0)}, \mathbf{Z}_{1,2}^{(i_0)}, \cdots, \mathbf{Z}_{1,\mathbf{k}}^{(i_0)})$ the population vector at time 1.

Next, after another unit of time, the type structure of the population grows from τ_{α}^{m} to τ_{α}^{m+1} . So, again in the induced population, we replace $\sigma_{\tau_{\alpha}}^{(1)}$ (m-1)-patterns rooted at individuals in the 1st generation with $\sigma_{\tau_{\alpha}}^{(2)}$ (m-1)-patterns

$$\left\{ (\tau_{\alpha}^{m+1})_{1}^{(2)}, (\tau_{\alpha}^{m+1})_{2}^{(2)}, \cdots, (\tau_{\alpha}^{m+1})_{\sigma_{\tau_{\alpha}}^{(2)}}^{(2)} \right\}$$

rooted at individuals in the 2nd generation in τ_{α}^{m+1} to obtain the population vector $\vec{\mathbf{Z}}_{2}^{(i_{0})} = (\mathbf{Z}_{2,1}^{(i_{0})}, \mathbf{Z}_{2,2}^{(i_{0})}, \cdots, \mathbf{Z}_{2,\mathbf{k}}^{(i_{0})})$ at time 2 and so on.

Induced Branching Process

Let $\vec{\mathbf{Z}}_{n}^{(i_0)} = (\mathbf{Z}_{n,1}^{(i_0)}, \mathbf{Z}_{n,2}^{(i_0)}, \cdots, \mathbf{Z}_{n,\mathbf{k}}^{(i_0)})$ be the population vector at time n, where $\mathbf{Z}_{n,i}^{(i_0)}$ is the number of the (m-1)-patterns of type α_i among the population

$$\left\{ (\tau_{\alpha}^{m+n-1})_{1}^{(n)}, (\tau_{\alpha}^{m+n-1})_{2}^{(n)}, \cdots, (\tau_{\alpha}^{m+n-1})_{\sigma_{\tau_{\alpha}}^{(n)}}^{(n)} \right\}$$

at time n.

Then such a process $\{\vec{\mathbf{Z}}_n\}_{n\geq 0}$ is called the *induced process* with the induced type set **A** from the random *m*-spread model \mathcal{R} .

From the construction, for almost every realization of the infinite spread pattern τ_{α} , there is a corresponding realization of the process $\{\vec{\mathbf{Z}}_n\}$ and, we directly obtain the following properties:

- (i) Each (m-1)-pattern rooted at a node in the *n*th level of τ_{α} represents an individual in the *n*th generation of $\{\vec{\mathbf{Z}}_n\}$.
- (ii) If $\alpha = \alpha_{i_0} \in \mathbf{A}$, then, for all $n = 0, 1, 2, \dots, |\vec{\mathbf{Z}}_n^{(i_0)}| = \sigma_{\tau_{\alpha}}^{(n)}$ with probability 1.

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(iii) The induced process $\{\vec{\mathbf{Z}}_n\}_{n\geq 0}$ from any random *m*-spread model \mathcal{R} is a multi-type Galton-Watson branching process with the type set \mathbf{A} and the offspring distribution $\{\mathbb{P}^{(i)}(\cdot)\}_{i=1}^{\mathbf{k}}$, where

$$\mathbb{P}^{(i_0)}(\vec{r}) = \sum_{\substack{\beta \in \mathcal{Q}_m \text{ s.t.} \\ \mathbf{T}(\beta) = \vec{r}}} f(\alpha_{i_0}, \beta)$$

and $\mathbf{T}(\beta) = (t_1, t_2, \cdots, t_{\mathbf{k}})$ is the vector with t_i as the number of the (m-1)-patterns of type α_i among $\left\{\beta_1^{(1)}, \beta_2^{(1)}, \cdots, \beta_{\sigma_{\beta}^{(1)}}^{(1)}\right\}, i = 1, 2, \cdots, \mathbf{k}.$

(iv) $\{\vec{\mathbf{Z}}_n\}_{n\geq 0}$ is positive regular.

We further assume that the induced branching process $\{\vec{\mathbf{Z}}_n\}_{n\geq 0}$ is non-singular.

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By the well-known limit theorems for classical multitype branching processes, we have the growth rate regarding the induced branching process:

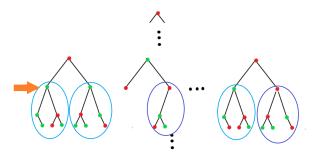
Lemma 3.

Let $\{\vec{\mathbf{Z}}_n\}_{n\geq 0}$ be the induced branching process from the random *m*-spread model \mathcal{R} . Let \mathbf{M} be the offspring mean matrix with the spectral radius $\rho > 1$ and the corresponding right and left eigenvectors $\vec{\mathbf{u}} = (\mathbf{u}_1, \cdots, \mathbf{u}_k)$ and $\vec{\mathbf{v}} = (\mathbf{v}_1, \cdots, \mathbf{v}_k)$ such that $\vec{\mathbf{v}} \cdot \vec{\mathbf{l}} = 1$ and $\vec{\mathbf{v}} \cdot \vec{\mathbf{u}} = 1$. Then, for every $i = 1, 2, \cdots, k$,

$$s_{\alpha_i}(\alpha_j) = \lim_{n \to \infty} s_{\alpha_i}^{(n)}(\alpha_j) \equiv \lim_{n \to \infty} \frac{\left| \prod(\mathbf{p}_{\alpha_i}^{(n)}) \right|_{\alpha_j}}{\sigma_{\alpha_i}^{(n)}} = \mathbf{v}_j \quad w.p.1.$$

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Induced Branching Process



$$s_{\alpha}(a_j; \{k_n\}_{n=1}^{\infty}) = \lim_{n \to \infty} \frac{\mathcal{O}_{a_j}(\tau_{\alpha}|_{\Delta_{s_n}^{s_n+1}(\tau_{\alpha})})}{\left|\Delta_{s_n}^{s_n+1}(\tau_{\alpha})\right|} = \sum_{\substack{i=1,2,\cdots,\mathbf{k} \text{ s.t.} \\ \alpha_i \in \theta(a_j)}} \mathbf{v}_i$$

References

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